

9.1

9. GROUP {SYMMETRY} {INVARIANCE} COVARIANCE MODELS

(S.A. ANDERSSON (COPENHAGEN) + MDP)

STUDY COVARIANCE STRUCTURES WHICH EXHIBIT SYMMETRIES INDUCED BY A (FINITE) GROUP OF (ORTHOGONAL) LINEAR TRANSFORMATIONS.

$X: p \times 1$ a random vector (\equiv MULTIVARIATE OBSERVATION)

$\Sigma: p \times p = \text{COV}(X) (\equiv \text{COVARIANCE MATRIX OF ERRORS})$

\mathcal{S}_p^+ = {all positive definite symmetric real $p \times p$ matrices}

$$[\mathcal{S}_p, \Sigma \in \mathcal{S}_p^+]$$

$G = \{I, g_1, \dots, g_{k-1}\}$ = Finite group of ORTHOGONAL MATRICES

$\mathcal{S}_G^+ = \{\Sigma \in \mathcal{S}_p^+ \mid g \Sigma g' = \Sigma \text{ for all } g \in G\}$

$[\mathcal{S}_G \text{ a REAL LINEAR SUBSPACE of } \mathcal{S}_p]$

INVERSE PROPERTY: $\Sigma \in \mathcal{S}_G^+ \Rightarrow \Sigma^{-1} \in \mathcal{S}_G^+$:

$$[g \Sigma = \Sigma g \text{ all } g \in G \Rightarrow \Sigma^{-1} g = g \Sigma^{-1} \text{ all } g \in G] \quad (g' = g^{-1})$$

(ALSO, $I \in \mathcal{S}_G^+$) \leftarrow NOT NEEDED - CAN ASSUME $g \Sigma_0 g' = \Sigma_0$ FOR A FIXED $\Sigma_0 \in \mathcal{S}_p^+$

GROUP SYMMETRY COVARIANCE HYPOTHESIS:

$$(*) \quad \boxed{\text{cov}(\tilde{X}) \equiv \Sigma \varepsilon \mathcal{J}_G^+}$$

This RESTRICTS THE FORM OF Σ .

IF TRUE, MORE ACCURATE INFERENCES CAN BE MADE.

WE WILL SHOW THAT THE MLE OF Σ UNDER (*) IS SIMPLE TO COMPUTE.

ALSO, CAN TEST ONE GROUP-INVARIANT MODEL* AGAINST ANOTHER

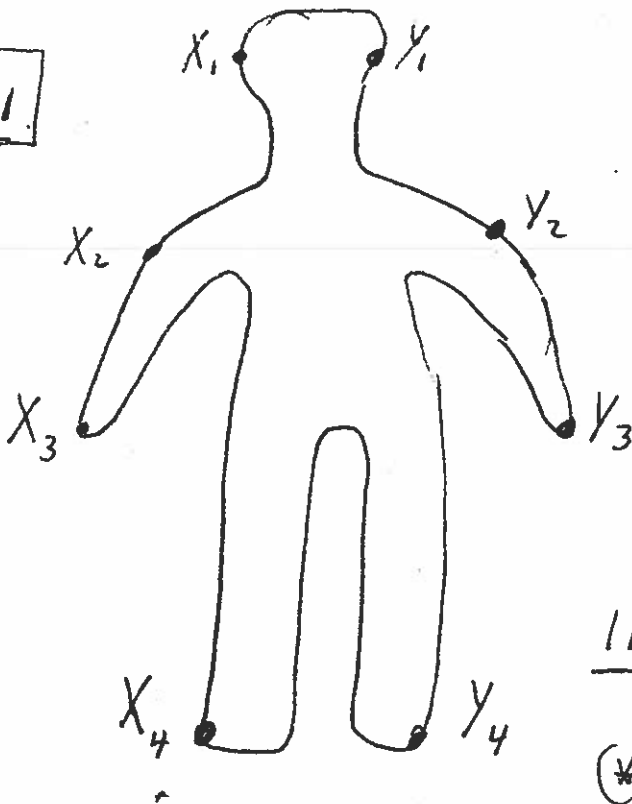
$$H_1: \Sigma \varepsilon \mathcal{J}_{G_1}^+ \quad \text{vs.} \quad H_2: \Sigma \varepsilon \mathcal{J}_{G_2}^+$$

$$\text{where } \mathcal{J}_{G_1}^+ \subseteq \mathcal{J}_{G_2}^+ \quad (G_2 \subseteq G_1)$$

WILL ALSO DISCUSS HOW TO RECOGNIZE WHEN A (LINEAR) RESTRICTION ON Σ ARISES FROM SOME GROUP G . (NOT ALL DO).

(FIRST, EXAMPLES.)

EX. 1



LEFT-RIGHT SYMMETRY

OBSERVE $\begin{pmatrix} X_1 \\ \vdots \\ X_4 \\ Y_1 \\ \vdots \\ Y_4 \end{pmatrix} = \begin{pmatrix} X \\ \sim \\ Y \end{pmatrix}$

IF ASSUME L-R SYMMETRY:

(*) $\text{cov} \begin{pmatrix} X \\ \sim \\ Y \end{pmatrix} = \text{cov} \begin{pmatrix} Y \\ \sim \\ X \end{pmatrix}$



(**) $\begin{matrix} \Sigma_{XX} & = & \Sigma_{YY} \\ \Sigma_{XY} & = & \Sigma_{YX} \end{matrix}$

UNRESTRICTED:

$$\Sigma = \begin{pmatrix} \Sigma_{XX} & \Sigma_{XY} \\ \Sigma_{YX} & \Sigma_{YY} \end{pmatrix}, \begin{matrix} \Sigma_{XX} = \Sigma'_{XX} \\ \Sigma_{YY} = \Sigma'_{YY} \\ \Sigma_{XY} = \Sigma_{YX} \end{matrix}$$

THIS RESTRICTION IS A GROUP SYMMETRY MODEL,

FOR $\begin{pmatrix} Y \\ \sim \\ X \end{pmatrix} = \begin{pmatrix} 0 & I \\ I & 0 \end{pmatrix} \begin{pmatrix} X \\ \sim \\ Y \end{pmatrix} = g \begin{pmatrix} X \\ \sim \\ Y \end{pmatrix}$

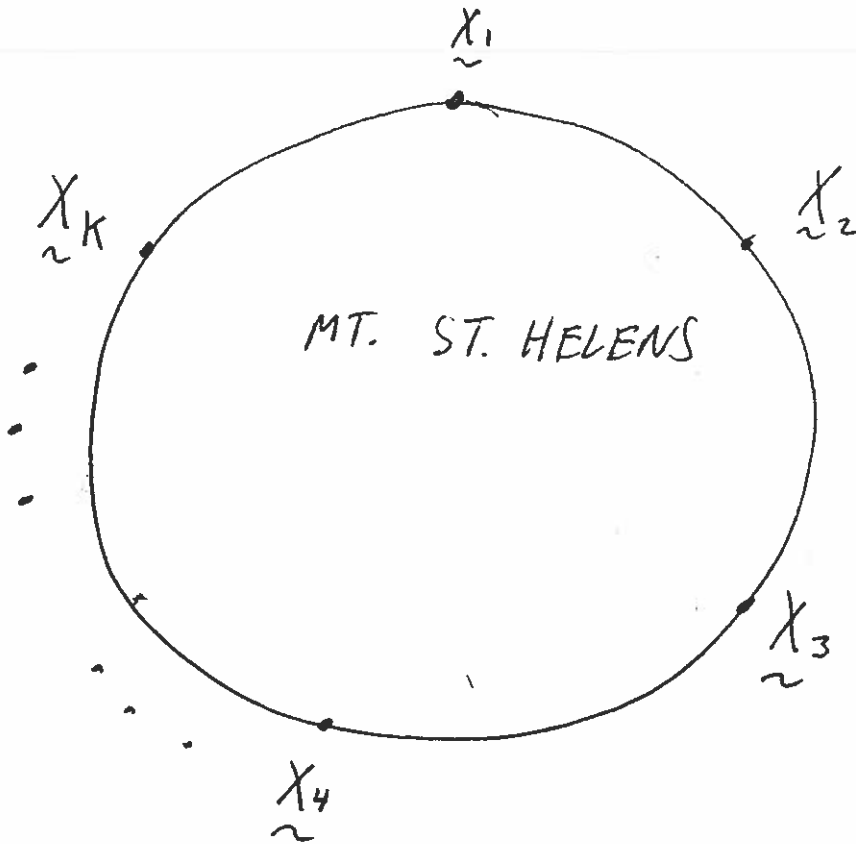
$\varphi(*) \Rightarrow \begin{pmatrix} \Sigma_{XX} & \Sigma_{XY} \\ \Sigma_{YX} & \Sigma_{YY} \end{pmatrix} = g \begin{pmatrix} \Sigma_{XX} & \Sigma_{XY} \\ \Sigma_{YX} & \Sigma_{YY} \end{pmatrix} g' = \begin{pmatrix} \Sigma_{YY} & \Sigma_{YX} \\ \Sigma_{XY} & \Sigma_{XX} \end{pmatrix}$

WHICH IS (**).

$G = \left\{ \begin{pmatrix} I & 0 \\ 0 & I \end{pmatrix}, \begin{pmatrix} 0 & I \\ I & 0 \end{pmatrix} \right\}, \mathcal{L}_G = \left\{ \begin{pmatrix} A & B \\ B & A \end{pmatrix} \mid \begin{matrix} A=A' \\ B=B' \end{matrix} \right\}$

EX. 2, 3, 4

K BLOCKS OF MULTIVARIATE OBSERV.



$$\tilde{X} = \begin{pmatrix} X_1 \\ \vdots \\ X_k \end{pmatrix}$$

$X_i: p \times 1 =$ (MULTIVARIATE) OBSERVATION AT i^{th} SEISMIC STATION.

EX. 2 CIRCULAR SYMMETRY:

$$(*) \text{COV} \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_k \end{pmatrix} = \text{COV} \begin{pmatrix} X_2 \\ X_3 \\ \vdots \\ X_1 \end{pmatrix} = \text{COV} \begin{pmatrix} X_3 \\ X_4 \\ \vdots \\ X_2 \end{pmatrix} = \dots = \text{COV} \begin{pmatrix} X_i \\ X_{i+1} \\ \vdots \\ X_{i+k-1} \end{pmatrix} \pmod{k}$$

$$\text{LET } P = \begin{pmatrix} 0 & I & 0 & \dots & 0 \\ 0 & 0 & I & & \vdots \\ \vdots & & \ddots & \ddots & I \\ I & \dots & & & 0 \end{pmatrix}$$

$$(*) \iff \text{cov}(\tilde{X}) = \text{cov}(P^n \tilde{X}), \quad n=0, 1, \dots, K-1$$

$n=1$ suffices: $[P \text{ generates the cyclic gp.}]$

$$\Sigma \equiv \text{cov}(\tilde{X}) = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} & \dots & \Sigma_{1K} \\ \Sigma_{21} & \Sigma_{22} & \dots & \Sigma_{2K} \\ \vdots & & \ddots & \vdots \\ \Sigma_{K1} & \dots & \dots & \Sigma_{KK} \end{pmatrix} \quad \text{SATISFIES } \Sigma = P \Sigma P'$$

UNDER CIRCULAR SYMMETRY.

$$\underline{K=3:} \quad \begin{pmatrix} 11 & 12 & 13 \\ 21 & 22 & 23 \\ 31 & 32 & 33 \end{pmatrix} = \begin{pmatrix} 22 & 23 & 21 \\ 32 & 33 & 31 \\ 12 & 13 & 11 \end{pmatrix}$$

$$\iff \begin{cases} \Sigma_{11} = \Sigma_{22} = \Sigma_{33} = A & (A=A^{-1}) \\ \Sigma_{12} = \Sigma_{23} = \Sigma_{31} = B & (\underline{\text{NOT}} B=B^{-1}) \end{cases}$$

$$\Leftrightarrow \Sigma = \begin{pmatrix} A & B & B' \\ B' & A & B \\ B & B' & A \end{pmatrix}, \quad \begin{matrix} A = A' \\ (B \neq B') \end{matrix}$$

$$\underline{k=4}: \quad \Sigma = \begin{pmatrix} A & B & C & B' \\ B' & A & B & C \\ C & B' & A & B \\ B & C & B' & A \end{pmatrix} \quad \begin{matrix} A = A' \\ (B \neq B') \\ C = C' \end{matrix}$$

(etc.)

These ARE GROUP SYMMETRY MODELS \mathcal{S}_G ,

$$G = \{I, P, \dots, P^{k-1}\} \cong \boxed{\text{cyclic group.}}$$

HERE DO NOT ASSUME $\text{COV}(\underset{\sim}{X}_i, \underset{\sim}{X}_j) = \text{COV}(\underset{\sim}{X}_j, \underset{\sim}{X}_i)$

IN MULTIVARIATE LITERATURE (OUTSIDE DENMARK),

"BLOCK CIRCULAR SYMMETRY MODEL" HAS REFERRED TO A DIFFERENT MODEL:

EX. 9

r=3:

$$\Sigma = \begin{pmatrix} A & B & B \\ B & A & B \\ B & B & A \end{pmatrix}$$

$$A = A^{-}$$

$$B = B^{-}$$

k=4:

$$\Sigma = \begin{pmatrix} A & B & C & B \\ B & A & B & C \\ C & B & A & B \\ B & C & B & A \end{pmatrix}$$

$$A = A^{-}$$

$$B = B^{-}$$

$$C = C^{-}$$

etc.

HERE, do ASSUME $\text{cov}(\tilde{x}_i, \tilde{x}_j) = \text{cov}(\tilde{x}_j, \tilde{x}_i)$.

(The difference ONLY shows up when the \tilde{x}_i ARE MULTIVARIATE)

This is a group symmetry model of G

$$G = \{I, P, \dots, P^{k-1}, Q, QP, \dots, QP^{k-1}\} \cong \boxed{\text{DIHEDRAL GROUP}}$$

$$Q = \begin{pmatrix} 0 & - & - & I \\ \vdots & & & I \\ & & & \vdots \\ I & - & - & 0 \end{pmatrix},$$

a REFLECTION.

CIRCULAR SYMMETRY
OR DIHEDRAL SYMMETRY

EX. 4 COMPLETE BLOCK SYMMETRY:

$$\Sigma = \begin{pmatrix} A & B & \dots & B \\ B & A & \dots & B \\ \vdots & \vdots & \ddots & \vdots \\ B & \dots & B & A \end{pmatrix} \quad \begin{array}{l} A = A^T \\ B = B^T \end{array}$$

This says

$$\begin{cases} \text{COV}(X_i) = \text{COV}(X_j) & \text{all } i, j \\ \text{COV}(X_i, X_j) = \text{COV}(X_l, X_m) & \text{all } i, j, l, m \end{cases}$$

$$\Leftrightarrow \Sigma = \Pi \Sigma \Pi^T, \quad \text{ALL (BLOCK) PERMUTATION MATRICES } \Pi.$$

GROUP SYMMETRY MODEL \mathcal{S}_G ,

$$G = \{ \text{all (block) permutation mats} \} \cong \boxed{\text{SYMMETRIC GR. ON } K \text{ LETTERS.}}$$

EX. 5

BUT WHAT ABOUT:

$$\Sigma = \begin{pmatrix} A & B \\ B' & A \end{pmatrix} \quad \begin{array}{l} A = A^{-1} \\ (B \neq B^{-1}) \end{array}$$

$$\Sigma = \begin{pmatrix} A & B & B \\ B' & A & B \\ B' & B' & A \end{pmatrix} \quad \begin{array}{l} A = A^{-1} \\ (B \neq B^{-1}) \end{array}$$

$$\Sigma = \begin{pmatrix} A & B & B & B \\ B' & A & B & B \\ B' & B' & A & B \\ B' & B' & B' & A \end{pmatrix} \quad \begin{array}{l} A = A^{-1} \\ (B \neq B^{-1}) \end{array}$$

etc.)

ARE THESE \mathcal{S}_G 's FOR SOME G ?

(They ARE linear subspaces of \mathcal{S}_{PK})

Which linear subspaces are \mathcal{S}_G 's?

(CLOSED UNDER INVERSES(?) FOR pos det's)

EX. 6 $\Sigma = \begin{pmatrix} \Sigma_{11} & \dots & \Sigma_{1k} \\ \vdots & & \vdots \\ \Sigma_{k1} & \dots & \Sigma_{kk} \end{pmatrix}, \quad \Sigma_{ij} = P_i \times P_j$

$$G = \left\{ \begin{pmatrix} \pm I_1 & 0 & \dots & 0 \\ 0 & \pm I_2 & & \vdots \\ \vdots & & \ddots & 0 \\ 0 & \dots & 0 & \pm I_k \end{pmatrix} \right\} \quad (2^k = \#G)$$

$$\mathcal{S}_G = \left\{ \Sigma = \begin{pmatrix} \Sigma_{11} & 0 & \dots & 0 \\ 0 & \Sigma_{22} & & 0 \\ \vdots & & \ddots & \\ 0 & \dots & 0 & \Sigma_{kk} \end{pmatrix} \right\} \quad \text{A GROUP SYMMETRY MODEL}$$

$\Sigma \in \mathcal{S}_G$ IS CLASSICAL HYPOTHESIS OF INDEPENDENCE OF k SETS OF VARIABLES

(See Example 6.7) (*TWA Ch. 9) NOTE: COMBINE EX. 6 + EX. 4
TO GET MODEL: $\Sigma_1 = \dots = \Sigma_k$
(ASSUME $P_1 = \dots = P_k$)

EX. 7 $\Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} \quad \Sigma_{ij} = P \times P$

$G =$ group generated by $\begin{pmatrix} 0 & -I \\ I & 0 \end{pmatrix} \equiv J$

$$= \{ I, J, -I, -J \}$$

$$\mathcal{S}_G = \{ \Sigma \mid J \Sigma J' = \Sigma \}$$

$$\begin{pmatrix} 0 & -I \\ I & 0 \end{pmatrix} \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} \begin{pmatrix} 0 & I \\ -I & 0 \end{pmatrix} = \begin{pmatrix} \Sigma_{22} & -\Sigma_{21} \\ -\Sigma_{12} & \Sigma_{11} \end{pmatrix}$$

$$J \Sigma J^{-1}$$

$$\therefore \Sigma \in \mathcal{S}_G \Leftrightarrow \left\{ \begin{array}{l} \Sigma_{11} = \Sigma_{22} (= \Sigma_{11}') \\ \Sigma_{12} = -\Sigma_{21} (= -\Sigma_{12}') \end{array} \right\}$$

$$\mathcal{S}_G = \left\{ \begin{pmatrix} A & -B \\ B & A \end{pmatrix} \mid \begin{array}{l} A = A^{-} \\ B = -B^{-} \end{array} \right\}$$

$$\xleftrightarrow{-1} A + iB \quad \text{COMPLEX HERMITIAN}$$

$$\text{HERE, } \Sigma \equiv \text{cov} \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \in \mathcal{S}_G$$

$$\Leftrightarrow \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} + i \begin{pmatrix} X_2 \\ X_1 \end{pmatrix} \sim \text{COMPLEX, } \text{MULTIVARIATE NORMAL DISTN.}$$

(ASSUMING NORMALITY)

NOTE: Let $\underline{z} = \underline{x}_1 + i\underline{x}_2$. Then $\text{cov}(i\underline{z}) = i(\text{cov } \underline{z})i^* = (ii^*)(\text{cov } \underline{z}) = \text{cov}(\underline{z})$.

\therefore IF $\underline{z} \sim$ Complex Normal (say $E \underline{z} = 0$ for simplicity), must require that $\underline{z} \sim i\underline{z}$, that is, $\underline{x}_1 + i\underline{x}_2 \sim i(\underline{x}_1 + i\underline{x}_2) = (-\underline{x}_2) + i\underline{x}_1$,

so require $(X_1, X_2) \sim (-X_2, X_1) \therefore \left\{ \begin{array}{l} \text{cov}(X_1) = \text{cov}(X_2) \equiv A \\ \text{cov}(X_1, X_2) = \text{cov}(-X_2, X_1) \equiv B \end{array} \right\}$ so $\text{cov} \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} = \begin{pmatrix} A & -B \\ B & A \end{pmatrix}$

EACH OF THESE EXAMPLES OF COVARIANCE MODELS HAVE BEEN ANALYZED SEPARATELY IN THE LITERATURE.

(i.e., MLE'S $\hat{\Sigma}$ FOUND UNDER SYMMETRY RESTRICTION
 + DISTRN OF MLE
 + LRT FOR TESTING ONE MODEL VS. ANOTHER
 etc.)

HOWEVER, DAWES AT COPENHAGEN HAVE WORKED OUT THEORY FOR ALL GROUP SYMMETRY MODELS.
 (ANDERSSON, BRONS, JENSEN)

TO ILLUSTRATE = FIND MLE $\hat{\Sigma}$
 UNDER RESTRICTION $\Sigma \in \mathcal{S}_G^+$.
 (ASSUME NORMALITY)

LET $S \sim W_p(m, \Sigma)$ (WISHART
 RAND. MTX)

$$f_{\Sigma}(S) = \text{CONST.} \frac{|S|^{\frac{m-p-1}{2}}}{|\Sigma|^{\frac{m}{2}}} e^{-\frac{1}{2} \text{tr} \Sigma^{-1} S}$$

$$\log f_{\Sigma}(S) = \text{CONST}^* - \frac{m}{2} \log |\Sigma| - \frac{1}{2} \text{tr} \Sigma^{-1} S$$

$$\text{Let } \Lambda \in \Sigma^{-1}$$

$$\text{RECALL } \Lambda \in \mathcal{S}_G^+ \iff \Sigma \in \mathcal{S}_G^+$$

$$\therefore \text{MAX}_{\Lambda \in \mathcal{S}_G^+} [\log |\Lambda| - m \text{tr} \Lambda S] = \textcircled{?}$$

$$\text{FACT: } \Lambda \in \mathcal{S}_G^+ \implies \text{tr} \Lambda S = \text{tr} g \Lambda g' S \quad (g \in G)$$

$$= \text{tr} \Lambda g' S g$$

$$= \text{tr} \Lambda \left[\frac{1}{\#(G)} \sum_{g \in G} g' S g \right]$$

$$= \text{tr} \Lambda \textcircled{S_G}$$

$$\text{FACT: } S_G \in \mathcal{S}_G^+$$

$$\therefore \text{MAX}_{\Lambda \in \mathcal{S}_G^+} [\log |\Lambda| - m \text{tr} \Lambda S]$$

$$= \text{MAX}_{\Lambda \in \mathcal{S}_G^+} [\log |\Lambda| - m \text{tr} \Lambda S_G]$$

$$\text{tr}(S - S_G) \Lambda = 0 \quad \forall \Lambda \in \mathcal{S}_G^+$$

$$S - S_G \perp \mathcal{S}_G^+$$

$$\downarrow$$

$$S_G = \text{LSE}$$

(= ORTHOG. PROJ
nt S onto \mathcal{S}_G^+)

BUT UNRESTRICTED MAX

OCCURS AT $\hat{\Lambda} = m S_G^{-1} \in S_G^+$

\therefore RESTRICTED MAX OCCURS THERE TOO.

\therefore RESTRICTED MLE $\hat{\Sigma} = \frac{1}{m} S_G$

(DANES): WISHART DIST'N CAN BE
DEFINED ON S_G^+ IN NATURAL WAY
 $\rightarrow S_G$ FOLLOWS THIS DISTN

FACT: $S_G =$ ORTHOGONAL PROJECTION
(\equiv LEAST SQUARES PROJ.)

OF S ONTO S_G (W.R.T.D $\langle A, B \rangle = \text{tr } AB$)

\therefore (FOR Σ , AS FOR MEANS):

MLE \equiv LEAST SQUARES EST

[GIVES ANOVA-LIKE DECOMPOSITION
FOR S when TESTING $\Sigma \in S_{G_1}$ vs S_{G_2} - see p. 9.19]

COMPUTATION OF $\sum^1 \equiv \frac{1}{m} \int_G$ EASY

IN EX. 1: $S \equiv \begin{pmatrix} S_{11} & S_{12} \\ S_{21} & S_{22} \end{pmatrix}$ $[S_{ij} = p \times p]$

$G = \left\{ \begin{pmatrix} I & 0 \\ 0 & I \end{pmatrix}, \begin{pmatrix} 0 & I \\ I & 0 \end{pmatrix} \right\}$, $\int_G = \left\{ \begin{pmatrix} A & B \\ B & A \end{pmatrix} \mid \begin{matrix} A=A' \\ B=B' \end{matrix} \right\}$

$\int_G = \frac{\begin{pmatrix} S_{11} & S_{12} \\ S_{21} & S_{22} \end{pmatrix} + \begin{pmatrix} S_{22} & S_{21} \\ S_{12} & S_{11} \end{pmatrix}}{2} = \begin{pmatrix} \frac{S_{11}+S_{22}}{2} & \frac{S_{12}+S_{21}}{2} \\ \frac{S_{12}+S_{21}}{2} & \frac{S_{11}+S_{22}}{2} \end{pmatrix}$

IN EX. 2:

$S = \begin{pmatrix} S_{11} & S_{12} & S_{13} & \dots & S_{1k} \\ S_{21} & S_{22} & S_{23} & & S_{2k} \\ \vdots & & \cdot & \cdot & \vdots \\ S_{k1} & \dots & & & S_{kk} \end{pmatrix}$ $[S_{ij} = p \times p]$

$G = \{ I, P, P^2, \dots, P^{k-1} \}$, $\int_G = \left\{ \begin{pmatrix} A & B & C & \dots & B' \\ B' & A & B & \dots & C \\ C' & B' & A & \dots & B \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ B & \dots & \dots & \dots & A \end{pmatrix} \right\}$

$$S_G = \frac{1}{K} \sum_{r=0}^{K-1} P^r S(P^r)^{-1}$$

$$= \frac{1}{K} \begin{pmatrix} S_{11} + z_2 z_2 & S_{12} + z_2 z_3 & S_{13} + z_2 z_4 & \dots \\ + \dots + K K & + \dots + K 1 & + \dots + K 2 & \dots \\ S_{21} + z_3 z_2 & \searrow & \searrow & \searrow \\ + \dots + 1 K & & & \\ \vdots & \searrow & \searrow & \searrow \\ \vdots & & & \searrow \\ & & & \searrow \end{pmatrix}$$

IN EX. 3:

$$S_G = \frac{1}{2K} \begin{pmatrix} 2 \begin{pmatrix} S_{11} + z_2 z_2 \\ + \dots + K K \end{pmatrix} & S_{12} + z_2 z_1 & S_{13} + z_3 z_1 & \dots \\ + z_3 z_3 + z_2 z_2 & + z_3 z_2 + z_2 z_3 & + z_4 z_2 + z_2 z_4 & \dots \\ + \dots & + K 1 + 1 K & + K 2 + 2 K & \dots \\ \searrow & \searrow & \searrow & \searrow \\ \searrow & \searrow & \searrow & \searrow \\ \searrow & \searrow & \searrow & \searrow \\ \searrow & \searrow & \searrow & \searrow \end{pmatrix}$$

IN EX. 4

$$S_G = \begin{pmatrix} \overset{1}{A} & \overset{1}{B} & \overset{1}{B} & \dots & \overset{1}{B} \\ \overset{1}{B} & \overset{1}{A} & \overset{1}{B} & & \overset{1}{B} \\ \overset{1}{B} & \overset{1}{B} & \overset{1}{A} & & \vdots \\ \vdots & & & \searrow & \overset{1}{B} \\ \overset{1}{B} & \dots & \overset{1}{B} & & \overset{1}{A} \end{pmatrix}$$

$$\overset{1}{A} = \frac{1}{K} [S_{11} + S_{22} + \dots + S_{KK}]$$

$$\overset{1}{B} = \frac{1}{K(K-1)} \left[\sum_{i \neq j} S_{ij} \right]$$

IN EX. 6:
$$\Sigma_G = \begin{pmatrix} S_{11} & 0 & \dots & 0 \\ 0 & S_{22} & & \vdots \\ \vdots & & \ddots & 0 \\ 0 & \dots & 0 & S_{kk} \end{pmatrix}$$

9.18

IN EX. 7:
$$\Sigma = \begin{pmatrix} S_{11} & S_{12} \\ S_{21} & S_{22} \end{pmatrix} \quad [S_{ij} : p \times p]$$

$$G = \{I; J\} \vee \{-I, -J\}, \quad J = \begin{pmatrix} 0 & -I \\ I & 0 \end{pmatrix}$$

$$\Sigma_G = \left\{ \begin{pmatrix} A & -B \\ B & A \end{pmatrix} \mid \begin{array}{l} A = A^* \\ B = -B^* \end{array} \right\}$$

$$\Sigma_G = \frac{\begin{pmatrix} S_{11} & S_{12} \\ S_{21} & S_{22} \end{pmatrix} + \begin{pmatrix} S_{22} & -S_{21} \\ -S_{12} & S_{11} \end{pmatrix}}{2} = \begin{pmatrix} \frac{S_{11} + S_{22}}{2} & \frac{S_{12} - S_{21}}{2} \\ \frac{S_{21} - S_{12}}{2} & \frac{S_{11} + S_{22}}{2} \end{pmatrix}$$

NOTE:
$$\left(\frac{S_{11} + S_{22}}{2} \right) + i \left(\frac{S_{12} - S_{21}}{2} \right)$$

$$\sim \text{COMPLEX WISHART} \left(\Sigma = \left(\frac{S_{11} + S_{22}}{2} \right) + i \left(\frac{S_{12} - S_{21}}{2} \right) \right)$$

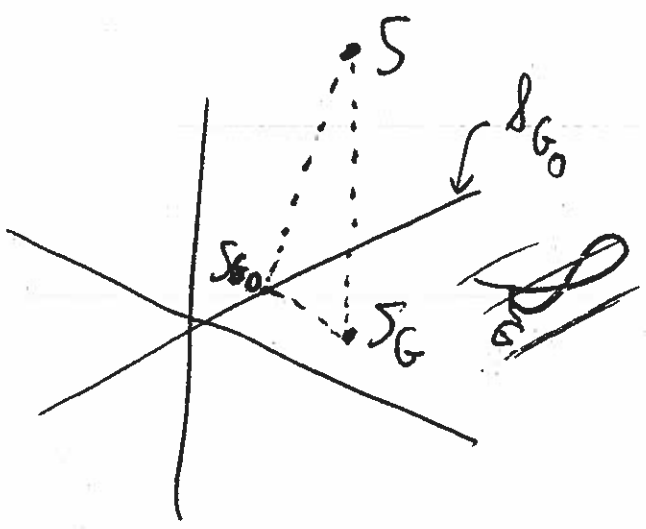
TESTING $H_0: \Sigma = S_{G_0}^+$ vs. $H_1: \Sigma \in S_G^+$ (BRIEF)

$$\left[\begin{array}{l} G_0 \supseteq G \\ S_{G_0}^+ \subseteq S_G^+ \end{array} \right]$$

"ANOVA" - DECOMPOSITION:

$$S \sim W_p(m, \Sigma) \quad (m \geq p)$$

$\frac{1}{m} S_G = \text{MLE of } \Sigma \text{ under } H$
 $= \text{ORTHOG PROJ OF } S \text{ ONTO } S_G$



$$S = S_{G_0} + (S_G - S_{G_0}) + (S - S_G)$$

"RESIDUAL"

① ANOVA-LIKE. HOWEVER (UNDER NORMALITY) (INVARIANT) TESTS OF H_0 vs. H BASED ON EIGENVALUES OF

$$(S_G - S_{G_0}) S_{G_0}^{-1}, \text{ NOT } (S_G - S_{G_0})(S - S_G)^{-1}$$

→ (NOT ROBUST) (EXAMPLE)

② LRT STATISTIC is $\det S_G / \det S_{G_0} \sim \prod \text{BETAS}$ UNDER H_0 | EXAMPLE IN §6.

Testing one Group symmetry model against another - DETAILS.

Consider two ^{finite} subgroups $G_0 \supseteq G$ of $\mathcal{O}_p = \{p \times p \text{ orthogonal matrices}\}$.

$$G_0 \supseteq G \Rightarrow \mathcal{S}_{G_0} \subseteq \mathcal{S}_G \quad (G \text{ imposes fewer restrictions})$$

\therefore We can test $H_0: \Sigma \in \mathcal{S}_{G_0}^+$ vs $H: \Sigma \in \mathcal{S}_G^+$, based on $J \sim N_p(n, \Sigma)$.

The LRT statistic is:

$$\lambda = \frac{\max_{\Sigma \in \mathcal{S}_{G_0}^+} \frac{|S|^{n-p-1}}{|S_0|^{n/2}} e^{-\frac{1}{2} \text{tr} \Sigma_0^{-1} S_{G_0}}}{\max_{\Sigma \in \mathcal{S}_G^+} \frac{|S|^{n-p-1}}{|S|^{n/2}} e^{-\frac{1}{2} \text{tr} \Sigma^{-1} S_G}} \quad (\text{see p. 13})$$

$$= \left(\frac{|S|}{|S_0|} \right)^{n/2} \quad (\text{see p. 14})$$

$$= \left(\frac{|S_G|}{|S_{G_0}|} \right)^{n/2}$$

$$S_0 = \frac{1}{\#(G_0)} \sum_{g_0 \in G_0} g_0 S g_0'$$

$$S_G = \frac{1}{\#(G)} \sum_{g \in G} g S g'$$

$$\hat{\Sigma}_0 = \frac{1}{n} S_{G_0}$$

$$\hat{\Sigma} = \frac{1}{n} S_G$$

Note that $0 \leq \lambda \leq 1$. The null distribution (under H_0) of λ can be approximated by a χ^2 -distribution: $-2 \log \lambda \sim \chi_{\dim H - \dim H_0}^2$. This approximation can be improved by the Box approximation (cf. T.W. Anderson's multivariate book). This requires knowledge of the moments of λ , equivalently, the moments of $\frac{|S_G|}{|S_{G_0}|}$. These moments are obtained by the following interesting identity:

$$(1) \quad E \left(\frac{|S_G|}{|S_{G_0}|} \right)^\alpha = \frac{E |S_G|^{2\alpha}}{E |S_{G_0}|^{2\alpha}} \quad \text{for all } \alpha > 0.$$

Both S_G and S_{G_0} can be shown to have generalized Wishart distributions, so that both $|S_G|$ and $|S_{G_0}|$ are distributed as products of independent (central) chi-square variates under H_0 , from which their moments are readily obtained. (See Anderson & Madsen (1998), Ann. Statist., Appendix A.7.)

"Symmetry & Lattice Conditional Independence in a Multivariate Normal Distribution"

It remains to verify (1). This requires the following lemma.

Lemma. Under H_0 , λ and S_{G_0} are independent. more generally, replace λ by any fn. of $g(S_0, S_0^{-1})$

Proof: Suppose that $\Sigma = \Sigma_0 \in \mathcal{J}_{G_0}^+$. Then (see below) $\Sigma_0^{\frac{1}{2}} \in \mathcal{J}_{G_0}$, hence $\Sigma_0^{-\frac{1}{2}} \in \mathcal{J}_{G_0}^+$. Then

$$\begin{aligned} \lambda &= \frac{|S_G|}{|S_{G_0}|} = \frac{|\frac{1}{\#G} \sum_g g S g^{-1}|}{|\frac{1}{\#G_0} \sum_{g_0} g_0 S g_0^{-1}|} = \frac{|\frac{1}{\#G} \sum_g \Sigma_0^{\frac{1}{2}} g S g^{-1} \Sigma_0^{-\frac{1}{2}}|}{|\frac{1}{\#G_0} \sum_{g_0} \Sigma_0^{\frac{1}{2}} g_0 S g_0^{-1} \Sigma_0^{-\frac{1}{2}}|} \\ &= \frac{|\frac{1}{\#G} \sum_g g (\Sigma_0^{-\frac{1}{2}} S \Sigma_0^{-\frac{1}{2}}) g^{-1}|}{|\frac{1}{\#G_0} \sum_{g_0} g_0 (\Sigma_0^{-\frac{1}{2}} S \Sigma_0^{-\frac{1}{2}}) g_0^{-1}|} \quad \left(\begin{array}{l} \text{since } g \Sigma_0^{\frac{1}{2}} g^{-1} = \Sigma_0^{\frac{1}{2}} \\ g_0 \Sigma_0^{-\frac{1}{2}} g_0^{-1} = \Sigma_0^{-\frac{1}{2}} \end{array} \right) \end{aligned}$$

But $\Sigma_0^{-\frac{1}{2}} S \Sigma_0^{-\frac{1}{2}} \sim W_p(n, \Sigma)$. Thus the distribution of λ does not depend on $\Sigma_0 \in \mathcal{J}_{G_0}^+$, so λ is an ancillary statistic under H_0 .

Next, because the pdf of S under H_0 is given by (pp 13-14)

$$f_{\Sigma_0}(S) = \frac{|S|^{-\frac{n-p-1}{2}}}{|Z_0|^{np/2}} e^{-\frac{1}{2} \text{tr} \Sigma_0^{-1} S_{G_0}},$$

a linear exponential family with natural parameter Σ_0^{-1} , it follows from the theory of exponential families that S_{G_0} is a complete and sufficient statistic under H_0 . Therefore, by Basu's Lemma (cf. Lehmann, Testing Statistical Hypotheses, Chapter 5), λ and S_{G_0} are independent under H_0 .

To see that $\Sigma \in \mathcal{J}_G^+ \Rightarrow \Sigma^{\frac{1}{2}} \in \mathcal{J}_G^+$, let $\Sigma = P D_\Sigma P'$ be the spectral decomposition of Σ , so that $\lambda = (\lambda_1, \dots, \lambda_p)$ are the eigenvalues of Σ . Let $h(x)$ be any polynomial such that $h(\lambda_i) = \lambda_i^{\frac{1}{2}}$, $i=1, \dots, p$. Then

$$\begin{aligned} \Sigma^{\frac{1}{2}} &= P D_\Sigma^{\frac{1}{2}} P' = P \begin{pmatrix} \lambda_1^{\frac{1}{2}} & & 0 \\ & \ddots & \\ 0 & & \lambda_p^{\frac{1}{2}} \end{pmatrix} P' = P \begin{pmatrix} h(\lambda_1) & & 0 \\ & \ddots & \\ 0 & & h(\lambda_p) \end{pmatrix} P' = P h(D_\Sigma) P' \\ &= h(P D_\Sigma P') = h(\Sigma). \text{ But } \Sigma \in \mathcal{J}_G^+ \Rightarrow h(\Sigma) \in \mathcal{J}_G^+. \text{ QED.} \end{aligned}$$

(ALL PROOFS, DISTRIBUTION THY, ETC., BASED ON:)

FIRST CHARACTERIZATION OF S_G (ANDERSSON 1975)

SPLITTING THEOREM: S_0 a linear subspace in S_p .

Then $S_0 = S_G$ for some G iff \exists a change of basis (\equiv NONSINGULAR TRANSFORMATION) s.t.

S_0 consists of all MATRICES OF THE BLOCK DIAGONAL FORM

$$\left[\begin{array}{ccc} \begin{pmatrix} \Sigma_1 & 0 \\ \vdots & \vdots \\ 0 & \Sigma_1 \end{pmatrix} & & 0 \\ & \begin{pmatrix} \Sigma_2 & 0 \\ \vdots & \vdots \\ 0 & \Sigma_2 \end{pmatrix} & \\ & & \ddots \\ 0 & & \begin{pmatrix} \Sigma_t & 0 \\ \vdots & \vdots \\ 0 & \Sigma_t \end{pmatrix} \end{array} \right]$$

WHERE:

EACH Σ_i RANGES OVER

$$S_{P_i} \equiv \left\{ \text{all real symmetric matrices} \right\} \equiv \int_{\mathbb{R}}^{(P_i)} : P_i \times P_i$$

OR

$$\left\{ \begin{array}{l} \text{all real symmetric matrices} \\ \text{of complex structure} \\ \cdot \begin{pmatrix} A & -B \\ B & A \end{pmatrix}, \quad \begin{array}{l} A = A^T \\ B = -B^T \end{array} \end{array} \right\} \equiv \int_{\mathbb{C}}^{(P_i)} : 2q_i \times 2q_i$$

OR

$$\left\{ \begin{array}{l} \text{all real symmetric matrices} \\ \text{of quaternion structure} \\ (4 \times 4 \text{ BLOCK}) \end{array} \right\} \equiv \int_{\mathbb{Q}}^{(P_i)} : 4q_i \times 4q_i$$

PROOF = (VIA SCHUR'S LEMMA)

$$\begin{aligned} \text{i.e., } S_0 &= \left(\int_{\mathbb{R}}^{(P_1)} \otimes I_{m_1} \right) \oplus \dots \oplus \left(\int_{\mathbb{R}}^{(P_k)} \otimes I_{m_k} \right) \\ &\oplus \left(\int_{\mathbb{C}}^{(Q_1)} \otimes I_{m_1} \right) \oplus \dots \oplus \left(\int_{\mathbb{C}}^{(Q_l)} \otimes I_{m_l} \right) \\ &+ \left(\int_{\mathbb{Q}}^{(R_1)} \otimes I_{s_1} \right) \oplus \dots \oplus \left(\int_{\mathbb{Q}}^{(R_m)} \otimes I_{s_m} \right) \end{aligned}$$

Thus, UNDER NORMALITY, A GROUP SYMMETRY MODEL REDUCES TO A PRODUCT OF INDEPENDENT "SIMPLE" MODELS WHERE THE DISTRIBUTION THEORY IS "EASY"

BUT: SPLITTING THM. IS AN EXISTENCE THEOREM. (EXTRINSIC CHAR'N).

DON'T KNOW HOW TO FIND THE APPROPRIATE BASIS TRANSFORMATION, OR IF ONE EXISTS.

IT DOES NOT TELL US HOW TO RECOGNIZE IF A GIVEN SUBSPACE \mathcal{S}_0 IS AN \mathcal{S}_G FOR SOME G .

NEED AN INTRINSIC CHARACTERIZATION. ("REFLEXIVE" + SEE P. 9.27)

CONSIDER SOME EXAMPLES OF \mathcal{S}_0 's:

$\mathcal{S}_0 =$ ALL (REAL SYMMETRIC) MATRICES OF FOLLOWING FORMS:

<p>① $\begin{pmatrix} A & B \\ B' & A \end{pmatrix}$</p>	<p>$A = A'$ $(B \neq B')$</p>	<p>$\mathcal{S}_G?$</p> <p><u>NO</u></p> <p>(NOT "REFLEXIVE")</p>
<p>CONSIDER $\begin{pmatrix} I & B \\ B' & I \end{pmatrix} = S$. Then $S^{-1} = \begin{pmatrix} (I - BB')^{-1} & * \\ * & (I - B'B)^{-1} \end{pmatrix}$ \therefore Require $I - BB' = I - B'B$, & $BB' = B'B$</p>		<p>← (INVERSE FAILS) BUT NOT NECESSARY</p>

So

②
$$\begin{pmatrix} A & B & \dots & B \\ B^{-1} & A^{-1} & & \\ \vdots & \ddots & A & B \\ B^{-1} & & B^{-1} & A \end{pmatrix}$$

$A = A^{-1}$
 $(B \neq B^{-1})$

SG (?)

NO

NOT REFLEXIVE

(INVERSE FAILS BUT NOT NECESSARY)

③
$$\begin{pmatrix} a & a & a & a \\ a & b & b & b \\ a & b & c & c \\ a & b & c & d \end{pmatrix}$$

a, b, c, d
Reals

BETTER:
$$\begin{pmatrix} a & a & a & a \\ a & b & b & b \\ a & b & c & c \\ a & b & c & d \end{pmatrix}$$

(I.E. $I \neq I_0$)
OK

~~NO~~ Yes

$I \neq I_0$

BUT NOT NECESS!

④
$$\begin{pmatrix} a & b & 0 \\ b & a & 0 \\ 0 & 0 & a \end{pmatrix}$$

a, b
Reals

SPECIAL CASE OF $\begin{pmatrix} a & b & c \\ b & a & c \\ c & c & a \end{pmatrix}$ - see above

NO?

INVERSE PROPERTY FAILS

$a^{33} \neq a^{11}, a^{22}$
BUT IS THIS REFLEXIVE?

⑤
$$\begin{pmatrix} a & b & 0 \\ b & a & b \\ 0 & b & a \end{pmatrix}$$

a, b
Reals

NO?

$$\begin{pmatrix} a & b \\ b & a \end{pmatrix} - \begin{pmatrix} 0 & a^{-1} \\ a^{-1} & 0 \end{pmatrix} \begin{pmatrix} 0 & b \\ b & 0 \end{pmatrix}$$

$$\neq \begin{pmatrix} a_1 & b_1 \\ b_1 & a_1 \end{pmatrix}$$

(INVERSE FAILS)

6

$$\begin{pmatrix} a & b & | & c_1 & b \\ b & a & | & b & c_2 \\ \hline c_1 & b & | & a & b \\ b & c_2 & | & b & a \end{pmatrix}$$

a, b,
c₁, c₂
REALS

NO?

INVERSE
FAILS
(BUT IS THIS REFLEXIVE?)

7

$$\begin{pmatrix} a & b & b \\ b & c & d \\ b & d & c \end{pmatrix}$$

a, b,
c, d
REALS

YES

REFLEXIVE

$$G = \left\{ \begin{pmatrix} 1 & 0 & 0 \\ 0 & \pi \end{pmatrix} \right\} \left| \begin{matrix} \pi \text{ } 2 \times 2 \\ \text{PERM} \\ \text{MTX} \end{matrix} \right.$$

8

$$\begin{pmatrix} a & b & b \\ b & c & d \\ b & d & e \end{pmatrix}$$

a, b, c,
d, e
REAL

NO?

INVERSE
FAILS
(BUT IS THIS REFLEXIVE?)

SECOND (INTRINSIC) CHARACTERIZATION OF \mathcal{S}_0 :

(CLASSICAL, IN ANDERSSON (1975), 1-LINE PF):

$$\mathcal{S}_0 = \mathcal{S}_G \text{ for some } G \iff \mathcal{S}_0 \text{ REFLEXIVE}$$

FOR $\mathcal{S} \in \mathcal{S}_p$, $\mathcal{Y} \in \mathcal{Y}_p (= \text{all } p \times p \text{ NON SING. MATX})$

DEFINE:

$$\mathcal{Y}(\mathcal{S}) = \{ g \text{ NONSINGULAR MATX} \mid gS = Sg \text{ all } S \in \mathcal{S} \}$$

$$\mathcal{S}(\mathcal{Y}) = \{ S \text{ SYMMETRIC MATX} \mid gS = Sg \text{ all } g \in \mathcal{Y} \}$$

DEFINITION:

$$\mathcal{S}_0 \text{ is REFLEXIVE} \iff$$

$$\boxed{\mathcal{S}_0 = \mathcal{S}(\mathcal{Y}(\mathcal{S}_0))}$$

$$\left(\begin{array}{l} \text{so } \mathcal{S}_0 = \mathcal{S}_G \text{ where} \\ G = \mathcal{Y}(\mathcal{S}_0) \\ \text{(A group!)} \end{array} \right)$$

NOTE: ALWAYS $\mathcal{S}_0 \subseteq \mathcal{S}(\mathcal{Y}(\mathcal{S}_0))$

A SHORTCUT:

IF $S_0 = \text{LINEAR SPAN} \{S_1, S_2, \dots, S_n\}$,

Then $\mathcal{L}(S_0) = \mathcal{L}(\{S_1, \dots, S_n\})$

THUS IN EX. ①,

NEED * COMPUTE

$$\mathcal{L} \left(\begin{Bmatrix} A & 0 \\ 0 & A \end{Bmatrix} \mid A = A^{-1} \right)$$

$$\sim \mathcal{L} \left(\begin{pmatrix} 0 & B \\ B^{-1} & 0 \end{pmatrix} \mid (B \neq B^{-1}) \right)$$